

Imamo domnevo H_0 in domnevo H_1 .

Če je model parametriziran s $\theta \in \Theta$, lahko ničelno domnevo H_0 določimo s $\theta \in \Theta_0$.

Alternativno domnevo H_1 lahko določimo kar z $\Theta_1 = \Theta \setminus \Theta_0$.

Pri stopnji tveganja α konstruiramo preizkus, da velja:

$$P_0(H_0 \text{ zavrnjena}) := P(H_0 \text{ zavrnjena} \mid H_0 \text{ drži}) \stackrel{=}{\leq} \alpha$$

Moč preizkusa definiramo kot:

$$P_1(H_0 \text{ zavrnjena}) := P(H_0 \text{ zavrnjena} \mid H_1 \text{ drži})$$

Odlučitev o zavrnitvi H_0 sprejmemo, če preizkusna statistika pade v kritično območje K_α .

	H_0 drži	H_1 drži
H_0 zavrnemo	napaka tipa 1 $P(\dots) \leq \alpha$	✓ $P(\dots) = 1 - \beta$ <small>moč testu</small>
H_0 ne zavrnemo	✓ $P(\dots) \geq 1 - \alpha$	napaka tipa 2 $P(\dots) = \beta$

3) Opazujemo $X_1, \dots, X_n \sim N(\mu, \sigma^2)$ NEP. Vemo, da je $\sigma = 5$.
Preizkušamo $H_0: \mu = 100$.

a) Naj bo $H_1: \mu < 100$. Skonstruiraj preizkus stopnje tveganja α .

$\hat{\mu} = \bar{X}$ preizkusna statistika

$$K_\alpha = (-\infty, c)$$

$\bar{X} \in K_\alpha \Rightarrow$ Zavrremo H_0

$$P_0(\bar{X} < c) \stackrel{!}{=} \alpha$$

$$X_i \sim N(100, 5^2)$$

$$\text{var}(\bar{X}) = \frac{1}{n^2} \sum_{i=1}^n \text{var}(X_i) = \frac{25}{n}$$

$$\Rightarrow \bar{X} \sim N\left(100, \frac{25}{n}\right)$$

$$P(\bar{X} < c) = P\left(\frac{\bar{X} - 100}{\frac{5}{\sqrt{n}}} < \frac{c - 100}{\frac{5}{\sqrt{n}}}\right) = F_{N(0,1)}\left(\frac{c - 100}{\frac{5}{\sqrt{n}}}\right) = \alpha$$

$$\frac{c - 100}{\frac{5}{\sqrt{n}}} = F_{N(0,1)}^{-1}(\alpha)$$

$$c - 100 = \frac{5}{\sqrt{n}} F_{N(0,1)}^{-1}(\alpha)$$

$$c = \frac{5}{\sqrt{n}} F_{N(0,1)}^{-1}(\alpha) + 100$$

b) Naj bo $H_1: \mu \neq 100$. Skonstruiraj preizkus stopnje tveganja α .



$$K_\alpha = (-\infty, \mu_0 - d) \cup (\mu_0 + d, \infty)$$

$|\bar{X} - \mu_0| > d \Rightarrow$ Zavrնemo H_0

$$P(|\bar{X} - \mu_0| > d) \stackrel{=}{\leq} \alpha$$

$$2P(\bar{X} - \mu_0 > d) \leq \alpha$$

$$2P\left(\frac{\bar{X} - \mu_0}{\frac{\sigma}{\sqrt{n}}} > \frac{d}{\frac{\sigma}{\sqrt{n}}}\right) \leq \alpha$$

$$1 - F_{N(0,1)}\left(\frac{d}{\frac{\sigma}{\sqrt{n}}}\right) \leq \frac{\alpha}{2}$$

$$d \cdot \frac{\sqrt{n}}{\sigma} \geq F_{N(0,1)}^{-1}\left(1 - \frac{\alpha}{2}\right)$$

$$d \geq \frac{\sigma}{\sqrt{n}} F_{N(0,1)}^{-1}\left(1 - \frac{\alpha}{2}\right)$$

4) Imamo podatke $X_1, \dots, X_5 \sim N(1, \sigma^2)$ NEP o finančni naložbi. Z verjetnostjo vsaj $1 - \alpha = 0,95$ se želimo izogniti investiranju, ko je $\hat{\sigma}^2 \geq \tilde{\sigma}_0^2$ za nek znan $\tilde{\sigma}_0^2 > 0$. Sestavi algoritem za sprejetje odločitve.

$$\hat{\sigma}^2 = \frac{1}{5} \sum_{i=1}^5 (X_i - 1)^2$$

$$T_{\sigma^2} := \frac{\hat{\sigma}^2 \cdot 5}{\sigma^2} \sim \chi^2(5), \text{ \u010d\u0117e je } \text{var}(X_i) = \sigma^2$$

Odl\u00f3itev o investiranju bo najbolj tvegana, ko je $\sigma^2 = \tilde{\sigma}_0^2$:

$$P_{\sigma^2}(T_{\sigma^2} < c) \text{ padajo\u0107a v } \sigma^2$$

$\Rightarrow P_{\tilde{\sigma}_0^2}(T_{\tilde{\sigma}_0^2} < c)$ je na $[\tilde{\sigma}_0^2, \infty)$ maksimalna pri $\sigma = \tilde{\sigma}_0$

Investiramo, \u010d\u0117e bo $T_{\sigma^2} < c$.

$$\Leftrightarrow \sum (X_i - 1)^2 < \sigma_0^2 \cdot c$$

Ko je $\sigma^2 > \sigma_0^2$, bomo investirali z verjetnostjo največ α .

$$P_{\sigma_0^2}(T_{\sigma_0^2} < c) = \alpha$$

$$F_{\chi^2(5)}(c) = \alpha$$

$$c = F_{\chi^2(5)}^{-1}(\alpha)$$

Preizkus na podlagi razmerja verjetij:

Imamo:

$$\begin{array}{l} H_0 \text{ s } \Theta_0 \\ H_1 \text{ s } \Theta_1 = \Theta \setminus \Theta_0 \end{array}$$

Preizkus na podlagi razmerja verjetij temelji na statistiki:

$$\Lambda = \frac{\sup_{\Theta_0} L(\theta|X)}{\sup_{\Theta} L(\theta|X)}$$

Oziroma:

$$\Lambda' = \frac{\sup_{\Theta_0} L(\theta|X)}{\sup_{\Theta} L(\theta|X)}$$

Wilksov izrek:

$$2 \ln \Lambda \xrightarrow[n \rightarrow \infty]{d} \chi^2(\dim \Theta - \dim \Theta_0)$$

5) X ... število klicev v Kopru na dan $\sim P_0(\lambda_x)$
 Y ... število klicev v Celju na dan $\sim P_0(\lambda_y)$

$$n = 100$$

$$\bar{X} = 20$$

$$\bar{Y} = 22$$

$\alpha = 0,05$ stopnja

$$H_0: \lambda_x = \lambda_y$$

$$H_1: \lambda_x \neq \lambda_y$$

$$\Theta = [0, \infty)^2$$

$$\Theta_0 = [0, \infty)$$

$$P(X=k) = \frac{\lambda_x^k}{k!} e^{-\lambda_x}$$

$$P(Y=k) = \frac{\lambda_y^k}{k!} e^{-\lambda_y}$$

$$L(\lambda_x, \lambda_y | X, Y) = \prod_{i=1}^n \frac{\lambda_x^{x_i}}{x_i!} e^{-\lambda_x} \cdot \frac{\lambda_y^{y_i}}{y_i!} e^{-\lambda_y}$$

$$\ell(\lambda_x, \lambda_y | X, Y) = \ln L = \sum_{i=1}^n (-\lambda_x + \ln \frac{\lambda_x^{x_i}}{x_i!} - \lambda_y + \ln \frac{\lambda_y^{y_i}}{y_i!})$$

$$= -n\lambda_x - n\lambda_y + \sum_{i=1}^n x_i \ln \lambda_x - \ln x_i! + y_i \ln \lambda_y - \ln y_i!$$

i) $H_0: \lambda_x = \lambda_y = \lambda$

$$\ell = -2n\lambda + \sum_{i=1}^n (x_i + y_i) \ln \lambda$$

$$\frac{\partial \ell}{\partial \lambda} = -2n + \sum_{i=1}^n \frac{x_i + y_i}{\lambda} = 0$$

$$\Rightarrow 2n = \frac{1}{\lambda} \sum_{i=1}^n (x_i + y_i)$$

$$\Rightarrow \lambda = \frac{1}{2n} \sum_{i=1}^n (x_i + y_i) = \frac{1}{2} (\bar{X} + \bar{Y})$$

To je maksimum, ker je drugi odvod manjši od 0.

$$ii) \frac{\partial L}{\partial \lambda_x}, \frac{\partial L}{\partial \lambda_y} = \dots$$

⋮

$$\Rightarrow \lambda_x = \bar{X}$$
$$\lambda_y = \bar{Y}$$

$$\Lambda = \frac{\prod \frac{x_i^{n_i} e^{-\bar{x}} \cdot y_i^{m_i} e^{-\bar{y}}}{x_i! y_i!}}{\prod \frac{(\frac{1}{2}(\bar{x} + \bar{y}))^{n_i + m_i} e^{-(\bar{x} + \bar{y})}}{x_i! y_i!}} = \dots = \frac{\bar{x}^{n\bar{x}} \bar{y}^{m\bar{y}}}{(\frac{1}{2}(\bar{x} + \bar{y}))^{n\bar{x} + m\bar{y}}}$$

H_0 zavrnemo, če je $\Lambda > c$.

Wilksov izrek: $2 \ln \Lambda \xrightarrow[n \rightarrow \infty]{d} \chi^2(1)$

$$P(\overbrace{2 \ln \Lambda}^{\sim \chi^2(1)} > d) = \alpha$$

$$1 - P(2 \ln \Lambda < d) = \alpha$$

$$1 - \alpha = P(2 \ln \Lambda < d)$$

$$d = F_{\chi^2(1)}^{-1}(1 - \alpha) \approx 3,84$$

V našem primeru dobimo $2 \ln \Lambda \approx 9,5$.

Torej H_0 zavrnemo.

